

Assessing inequality using percentile shares

An application to Swiss tax data

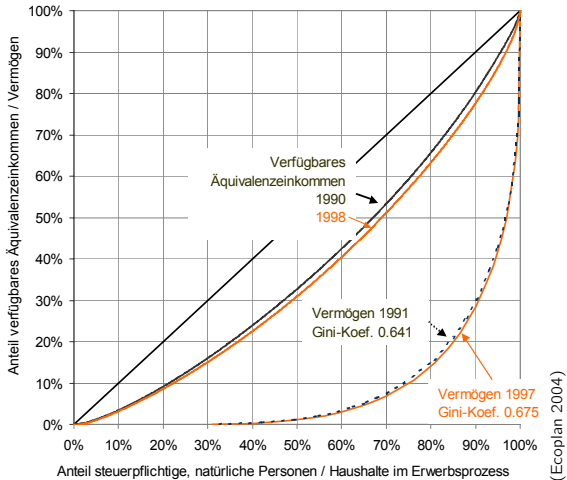
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HAPPY
BIRTHDAY





- ▷ http://www.youtube.com/watch?v=slTF_XXoKAQ
- ▷ https://www.ted.com/talks/dan_ariely_how_equal_do_we_want_the_world_to_be_you_d_be_surprised

Outline

- Motivation
- Estimation of percentile shares
- The `pshare` Stata command
- Examples
- Small sample bias

Estimation of percentile shares

- Outcome variable of interest, e.g. income: Y
- Distribution function: $F(y) = \Pr\{Y \leq y\}$
- Quantile function: $Q(p) = F^{-1}(p) = \inf\{y | F(y) \geq p\}$, $p \in [0, 1]$
- Lorenz ordinates:

$$L(p) = \int_{-\infty}^{Q_p} y dF(y) \Big/ \int_{-\infty}^{\infty} y dF(y)$$

- Finite population form:

$$L(p) = \sum_{i=1}^N y_i \mathcal{I}\{y_i \leq Q_p\} \Big/ \sum_{i=1}^N y_i$$

- Percentile share: proportion of total outcome within quantile interval $[Q_{p_{\ell-1}}, Q_{p_{\ell}}]$, $p_{\ell-1} \leq p_{\ell}$

$$S_{\ell} = L(p_{\ell}) - L(p_{\ell-1})$$

Estimation of percentile shares

- Estimation given sample of size n :

$$\widehat{S}_\ell = \widehat{L}(p_\ell) - \widehat{L}(p_{\ell-1})$$

$$\widehat{L}(p) = (1 - \gamma)\widetilde{Y}_{j-1} + \gamma\widetilde{Y}_j \quad \text{where } \widehat{p}_{j-1} < p \leq \widehat{p}_j \text{ with } \widehat{p}_j = \frac{j}{n}$$

$$\widetilde{Y}_j = \frac{\sum_{i=1}^j y_{(i)}}{\sum_{i=1}^n y_i} \quad \text{where } y_{(i)} \text{ refers to ordered values}$$

$$\gamma = \frac{p - \widehat{p}_{j-1}}{\widehat{p}_j - \widehat{p}_{j-1}} \quad \text{(linear interpolation)}$$

- Standard errors

- ▶ approximate standard errors can be obtained by the estimating equations approach as proposed by Binder and Kovacevic (1995)
- ▶ supports complex survey data and joint estimation across subpopulations or repeated measures
- ▶ alternative: bootstrap

Estimation of percentile shares: standard errors

- Let θ be a parameter interest and λ be a vector of nuisance parameters. Furthermore, let $u_\theta(y_i, \theta, \lambda)$ and $u_\lambda(y_i, \lambda)$ be estimating functions such that, in the (finite) population, θ and λ are the solutions to

$$U_\theta(\theta, \lambda) = \sum_{i=1}^N u_\theta(y_i, \theta, \lambda) = 0 \quad \text{and} \quad U_\lambda(\lambda) = \sum_{i=1}^N u_\lambda(y_i, \lambda) = 0$$

- Following Kovacević and Binder (1997), the sampling variance of $\hat{\theta}$ can be approximated by a variance estimate of

$$\sum_s w_s u^*(y_s, \hat{\theta}, \hat{\lambda})$$

where w_j are sampling weights and

$$u^*(y_i, \theta, \lambda) = \left(-u_\theta(y_i, \theta, \lambda) + \frac{\partial U_\theta}{\partial \lambda} \left[\frac{\partial U_\lambda}{\partial \lambda} \right]^{-1} u_\lambda(y_i, \lambda) \right) \left[\frac{\partial U_\theta}{\partial \theta} \right]^{-1}$$

Estimation of percentile shares: standard errors

- For percentile shares, $\theta = S$ and $\lambda = \begin{bmatrix} Q_1 \\ Q_2 \end{bmatrix}$.
- The estimating functions are:

$$u_\theta = y_i \mathcal{I}\{y_i \leq Q_2\} - y_i \mathcal{I}\{y_i \leq Q_1\} - y_i S$$

$$u_\lambda = \begin{bmatrix} \mathcal{I}\{y_i \leq Q_1\} - p_1 \\ \mathcal{I}\{y_i \leq Q_2\} - p_2 \end{bmatrix}$$

- Hence:

$$\begin{aligned} u^* &= \frac{y_i \mathcal{I}\{y_i \leq Q_2\} - y_i \mathcal{I}\{y_i \leq Q_1\} - y_i S}{\sum y_i} \\ &= \frac{(y_i - Q_2) \mathcal{I}\{y_i \leq Q_2\} - (y_i - Q_1) \mathcal{I}\{y_i \leq Q_1\} + Q_2 p_2 - Q_1 p_1 - y_i S}{\sum y_i} \end{aligned}$$

Estimation of percentile shares: some extensions

- Percentile share “density”:
 - ▶ particularly useful for graphing

$$D_\ell = \frac{S_\ell}{p_\ell - p_{\ell-1}} = \frac{L(p_\ell) - L(p_{\ell-1})}{p_\ell - p_{\ell-1}}$$

- Totals:

$$T_\ell = \sum_{i=1}^N y_i \mathcal{I}\{Q_{p_{\ell-1}} < y_i \leq Q_{p_\ell}\} = S_\ell \cdot \sum_{i=1}^N y_i$$

- Averages:
 - ▶ again, useful for graphing
 - ▶ useful if you are also interested in levels, not just distribution

$$A_\ell = \frac{T_\ell}{(p_\ell - p_{\ell-1}) \cdot N}$$

Estimation of percentile shares: some extensions

- Contrasts:

- ▶ useful for comparing distributions, e.g. changes over time
- ▶ standard errors easily computed using delta method

$$S_\ell^A - S_\ell^B \quad S_\ell^A / S_\ell^B \quad \ln(S_\ell^A / S_\ell^B) \quad \dots$$

- Renormalization (using a different total):

- ▶ useful, e.g., to analyze income components or subpopulation shares

$$L^*(p) = \sum_{i=1}^N y_i \mathcal{I}\{y_i \leq Q_p\} / T$$
$$S_\ell^* = L^*(p_\ell) - L^*(p_{\ell-1})$$

with T whatever you like it to be (e.g. the total of variable Z or the total across subpopulations)

Estimation of percentile shares: some extensions

- Concentration shares:
 - ▶ compute shares while ordering by a different variable
 - ▶ useful for analyzing relations between variables (wealth and income, pre- and post-tax income, etc.)

$$L^Z(p) = \frac{\sum_{i=1}^N y_i \mathcal{I}\{z_i \leq Q_p^Z\}}{\sum_{i=1}^N y_i}$$
$$S_\ell^Z = L^Z(p_\ell) - L^Z(p_{\ell-1})$$

- Often a combination of renormalization and using a different ordering variable is useful (e.g. to analyze redistribution).

The `pshare` Stata command

- `pshare estimate`
 - ▶ estimates the percentile shares and their variance matrix
 - ▶ arbitrary cutoffs for the percentile groups
 - ▶ joint estimation across multiple outcome variables or subpopulations
 - ▶ shares as proportions, densities, totals, or averages
 - ▶ etc.
- `pshare contrast`
 - ▶ computes contrasts between outcome variables or subpopulations
 - ▶ differences, ratios, or log ratios
- `pshare stack`
 - ▶ displays percentile shares as stacked bar chart
- `pshare histogram`
 - ▶ displays percentile shares as histogram

Example: quintile shares (the default)

```
. sysuse nlsw88
(NLSW, 1988 extract)
. pshare estimate wage, percent
Percentile shares (percent)      Number of obs   =      2,246
```

wage	Coef.	Std. Err.	[95% Conf. Interval]	
0-20	8.018458	.1403194	7.743288	8.293627
20-40	12.03655	.1723244	11.69862	12.37448
40-60	16.2757	.2068139	15.87013	16.68127
60-80	22.47824	.2485367	21.99085	22.96562
80-100	41.19106	.6246426	39.96612	42.41599

- top 20% percent of the population get 41% of wages
- bottom 20% only get 8% of wages, etc.

Example: bottom 50%, mid 40%, and top 10%

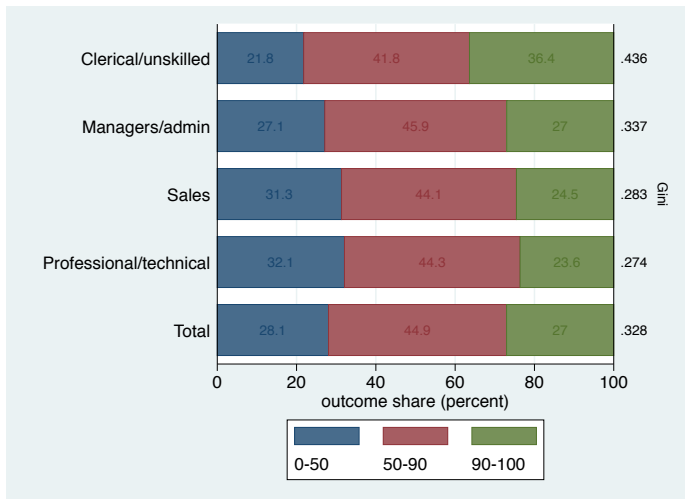
```
. pshare estimate wage, percent percentiles(50 90)
```

```
Percentile shares (percent)      Number of obs   =      2,246
```

wage	Coef.	Std. Err.	[95% Conf. Interval]	
0-50	27.59734	.3742279	26.86347	28.33121
50-90	45.86678	.4217771	45.03967	46.6939
90-100	26.53588	.682887	25.19672	27.87503

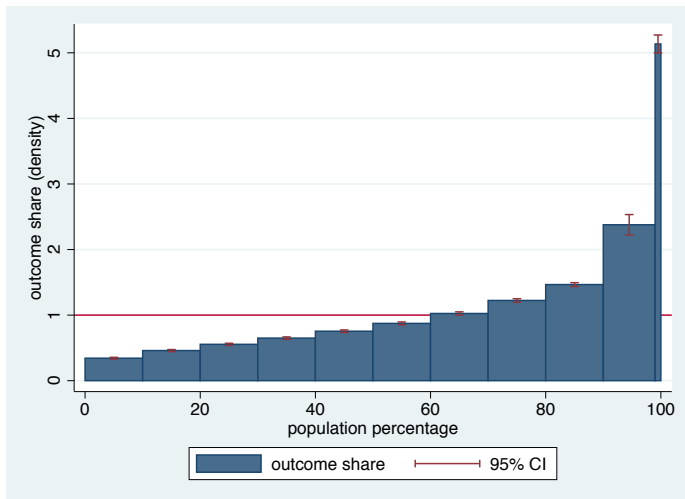
Example: stacked bars plot

- . pshare estimate wage if occ<=4, percent p(50 90) over(occ) total gini
(output omitted)
- . pshare stack, values sort(gini tlast descending)



Example: histogram of densities

- . pshare estimate wage, density percentiles(10(10)90 99)
(output omitted)
- . pshare histogram, yline(1)



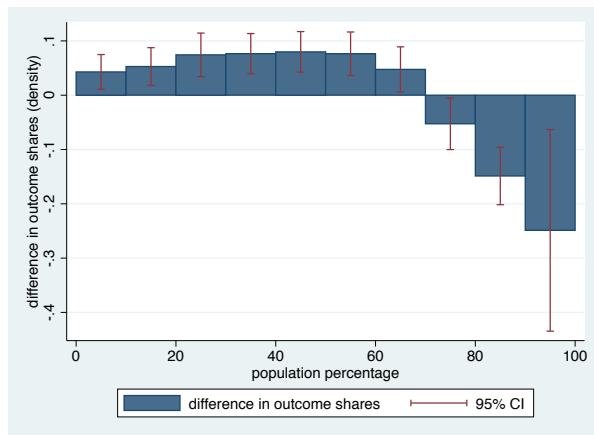
Example: histogram of densities

- Interpretation

- ▶ Take 100 dollars and divide them among 100 people who line up along the x-axis.
- ▶ The heights of the bars shows you how much each one gets.
- ▶ If all get the same, then everyone would get one dollar (red line).
- ▶ However, according to the observed distribution, the rightmost person would get five of the 100 dollars, the next 9 would get about two and a half dollars each, . . . , the bottom 10 only get 35 cents each.

Example: contrasts

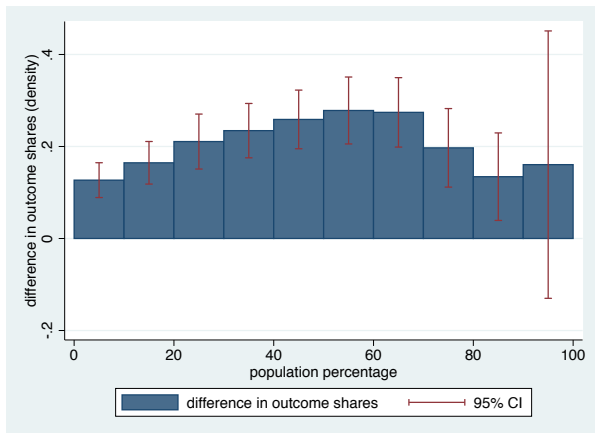
- . quietly pshare estimate wage, n(10) density over(union)
- . quietly pshare contrast 0
- . pshare histogram



- bottom 70% percent are *relatively* better off if unionized

Example: contrasts

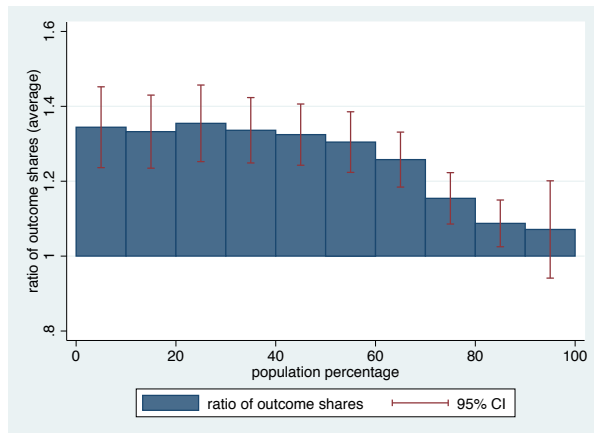
```
. pshare estimate wage, n(10) density over(union) contrast(0) normalize(0:) histogram  
(output omitted)
```



- everyone is *absolutely* better off if unionized (between about 15% and 25% of average nonunion wages)

Example: contrasts

```
. pshare estimate wage, n(10) average over(union) contrast(0, ratio) histogram  
(output omitted)
```



- bottom 50% of unionized are about 30% better off than bottom 50% of nonunionized; at the top the advantage shrinks to 10%

Example: concentration shares

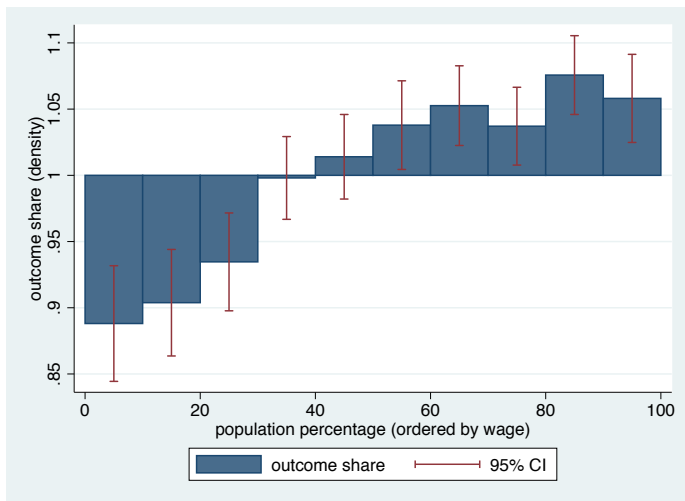
```
. pshare estimate hours, n(10) density pvar(wage)
Percentile shares (density)      Number of obs   =      2,242
```

hours	Coef.	Std. Err.	[95% Conf. Interval]	
0-10	.8880782	.0222773	.8443919	.9317646
10-20	.9038126	.0205245	.8635637	.9440616
20-30	.934641	.0188478	.8976801	.971602
30-40	.9980166	.0159431	.9667519	1.029281
40-50	1.014016	.0162895	.9820715	1.04596
50-60	1.037906	.0170757	1.00442	1.071392
60-70	1.052623	.0153487	1.022524	1.082722
70-80	1.037115	.0149871	1.007725	1.066505
80-90	1.075704	.0151754	1.045945	1.105464
90-100	1.058088	.0169731	1.024803	1.091372

(percentile groups with respect to wage)

```
. pshare histogram, base(1)
```

Example: concentration shares



- the 10% with the highest wages work 5.8% longer hours
- the 10% with the lowest wages work 11.2% shorter hours

Some examples with “real” data

- Tax data from canton of Bern, Switzerland, 2002 and 2012
- individual level data from personal tax forms
- information on income components, deductions, assets, etc.
- units of analysis in following examples are “tax units”

```
. describe
```

```
Contains data from BE-02-12.dta
```

```
obs:      1,153,709
```

```
vars:      10
```

```
28 Apr 2016 15:17
```

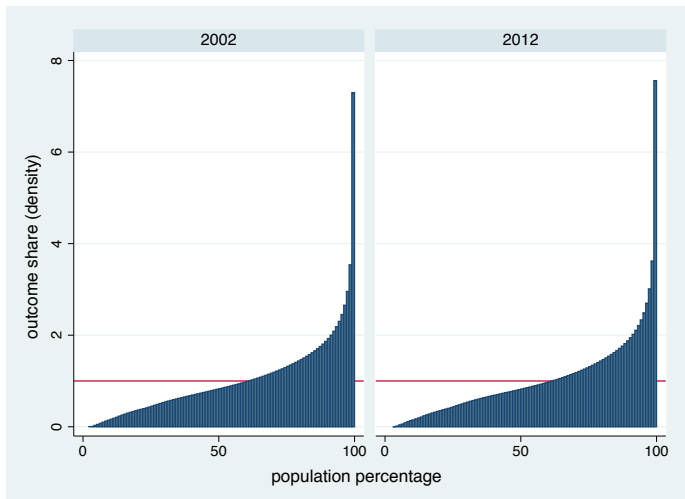
```
size:     48,455,778
```

variable name	storage type	display format	value label	variable label
year	int	%9.0g		Year
hhid	double	%10.0g		Household ID
earnings	float	%9.0g		Labor market income
capincome	float	%9.0g		Capital income
transfers	float	%9.0g		Transfer income
tax	float	%9.0g		Tax
heritage	long	%10.0gc		Received heritage
income	float	%9.0g		Total income
aftertax	float	%9.0g		After tax income
wealth	float	%9.0g		Net wealth

```
Sorted by:
```

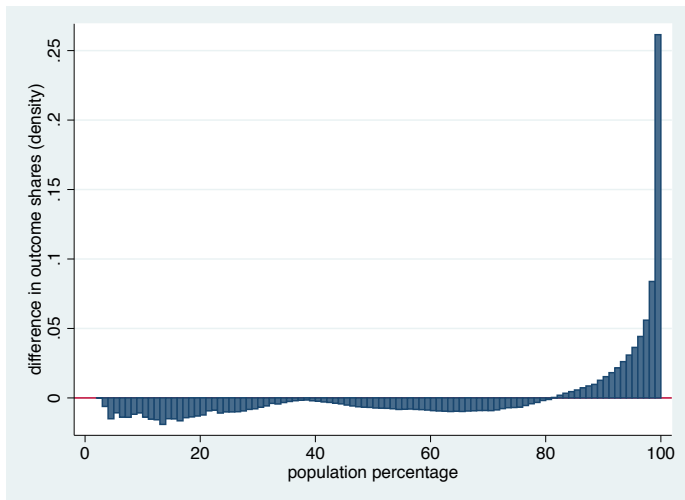
Distribution of total income in 2002 and 2012

- . pshare estimate income, n(100) nose density over(year)
(output omitted)
- . pshare histogram, yline(1)



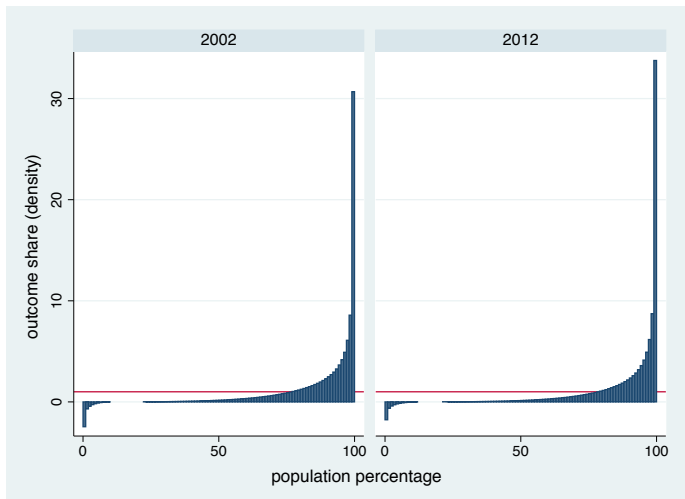
Change in income distribution from 2002 to 2012

- . pshare contrast
(output omitted)
- . pshare histogram, yline(0)



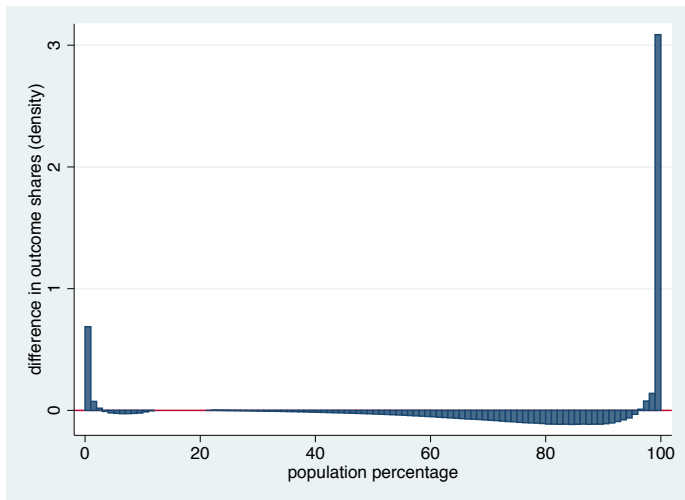
Distribution of net wealth in 2002 and 2012

- . pshare estimate wealth, n(100) nose density over(year)
(output omitted)
- . pshare histogram, yline(1)



Change in wealth distribution from 2002 to 2012

- . pshare contrast
(output omitted)
- . pshare histogram, yline(0)



Income composition by income percentiles (2012)

```
. keep if year==2012
(553,976 observations deleted)

. drop year

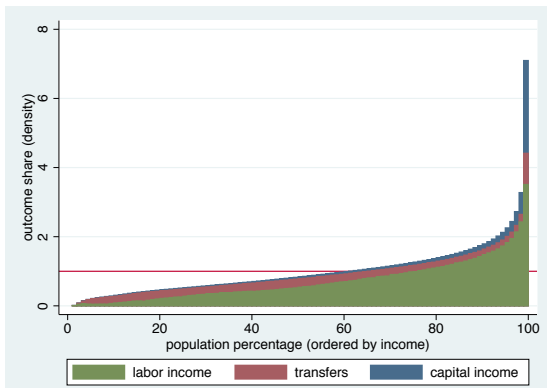
. drop if hhid>=.
(11,720 observations deleted)

. collapse (sum) earnings-wealth, by(hhid) fast // generate households

. generate earn_trans = earnings + transfers

. quietly pshare estimate income earn_trans earnings, n(100) nose density ///
> pvar(income) normalize(income)

. pshare histogram, overlay yline(1) fintensity(100) color(*.8) ///
> legend(order(3 "labor income" 2 "transfers" 1 "capital income") rows(1))
```



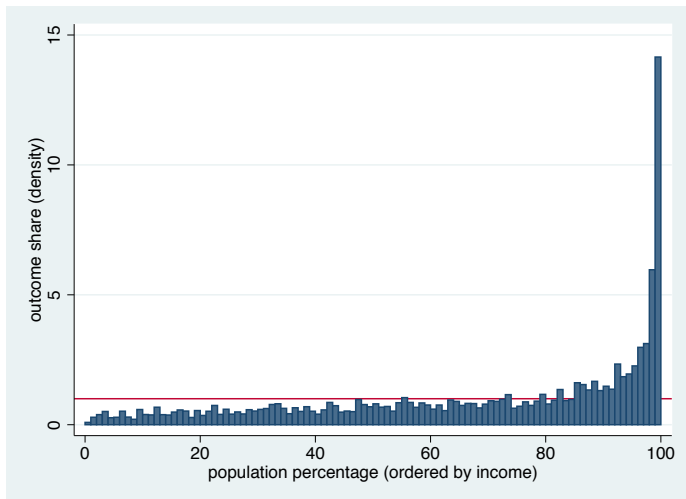
Income composition in relative terms (2012)

```
. generate earn_trans_cap = income
. quietly pshare estimate income earn_trans_cap earn_trans earnings, ///
>   p(10(1)99) nose density pvar(income) normalize(income)
. quietly pshare contrast income, ratio
. pshare histogram, overlay finten(100) color(*.8) base(0) ///
>   legend(order(3 "labor income" 2 "transfers" 1 "capital income") rows(1))
```



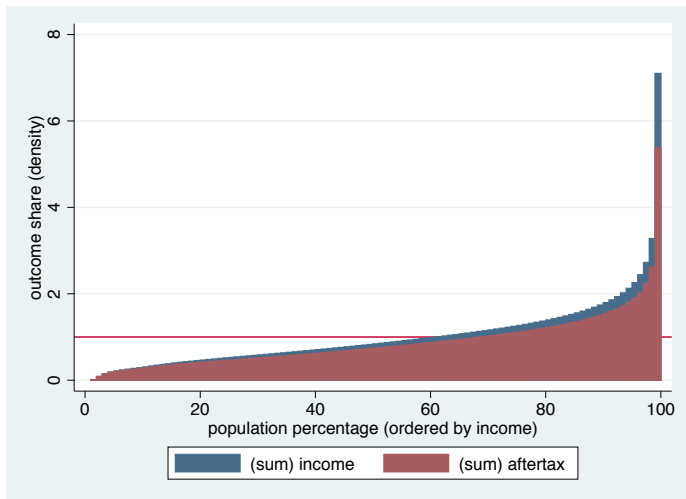
Received heritage by income percentiles (2012)

- . quietly pshare estimate heritage, n(100) nose density pvar(income)
(output omitted)
- . pshare histogram, yline(1)



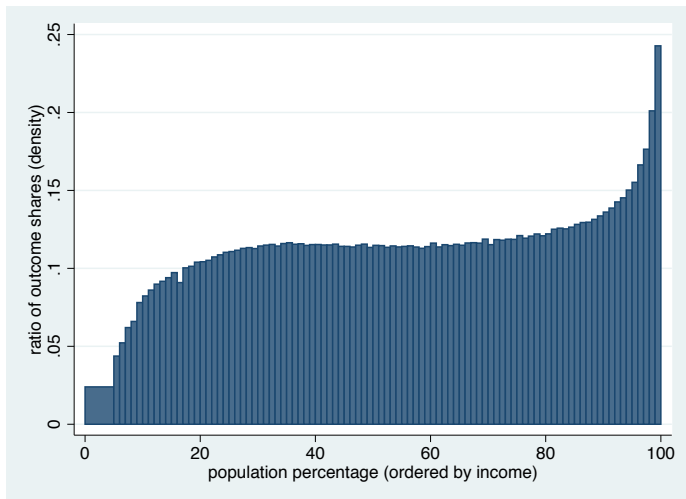
Pre-tax and post-tax income (2012)

```
. quietly pshare estimate income aftertax, n(100) nose density normalize(income)
> pvar(income)
(output omitted)
. pshare histogram, yline(1) overlay finten(100) color(*.8)
```



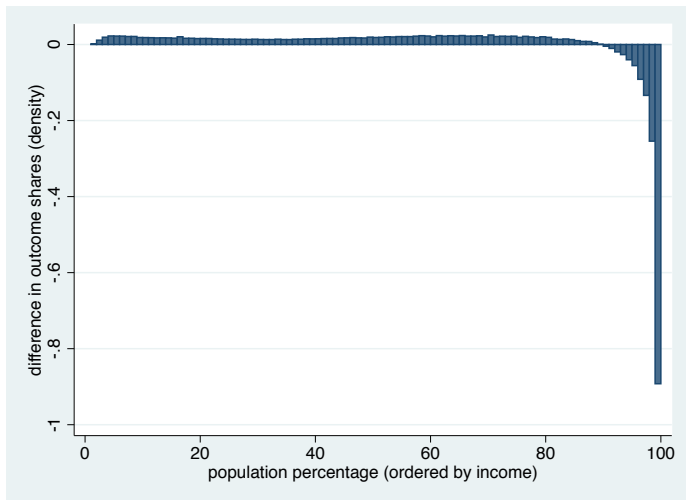
Tax rate by income percentiles (2012)

```
. quietly pshare estimate income tax, p(5(1)99) nose density ///  
>     normalize(income) pvar(income)  
  
. quietly pshare contrast income, ratio  
  
. pshare histogram, base(0) ylabel(0(.05).25)
```



“Winners” and “losers” from taxation (2012)

- . quietly pshare estimate income aftertax, n(100) nose density pvar(income)
- . quietly pshare contrast income
- . pshare histogram

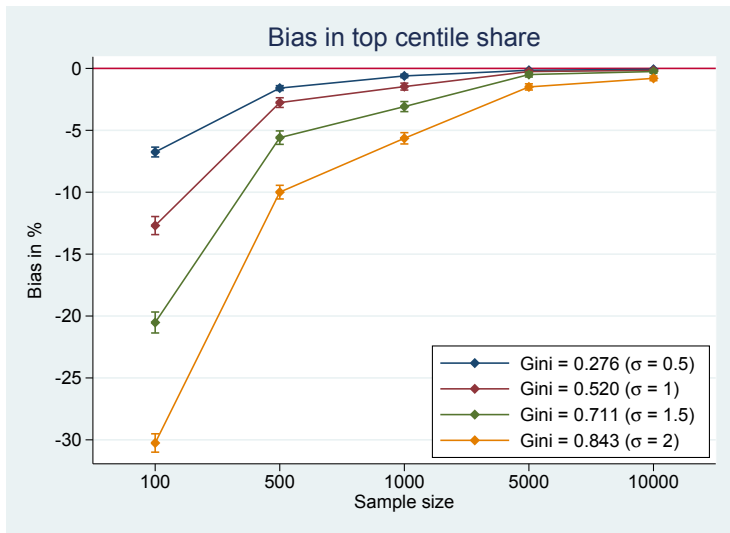


Small sample bias

- Percentile shares are affected by small sample bias.
- The top percentile share is typically underestimated.
- The problem is difficult to fix.
 - ▶ Corrections could be derived based on parametric assumptions.
 - ▶ Smoothing out the data by adding random noise can be an option, but this also requires parametric assumptions.
 - ▶ I evaluated a non-parametric small-sample correction using a bootstrap approach: the bias in bootstrap samples is used to derive correction factors for the main results.
 - ▶ This works very well in terms of removing bias (unless the distribution is extremely skewed).
 - ▶ **However:** MSE increases compared to uncorrected results!
 - ▶ Any ideas? Can Extreme Value Estimation be used to improve the estimates? Or would it be better to leave the point estimates as is and focus on obtaining bias-corrected CIs that have the correct size?

Small sample bias: how bad is the problem?

- Simulation: relative bias in top 1% share using a log-normal distribution



Software and paper

- Software:

```
. ssc install pshare
```

- Paper (forthcoming in the Stata Journal):

- ▶ Jann, Ben. 2015. Assessing inequality using percentile shares. University of Bern Social Sciences Working Papers No. 13. <https://ideas.repec.org/p/bss/wpaper/13.html>

Lorenz curves

- Should you still be attached to Lorenz curves/concentration curves, I wrote a companion command with similar functionality:

```
. ssc install lorenz
```

- Paper:
 - ▶ Jann, Ben. 2016. Estimating Lorenz and concentration curves in Stata. University of Bern Social Sciences Working Papers No. 15. <https://ideas.repec.org/p/bss/wpaper/15.html>

References

- Ecoplan (2004). Verteilung des Wohlstands in der Schweiz. Bern: Eidgenössische Steuerverwaltung.
- Binder, D. A., M. S. Kovacevic (1995). Estimating Some Measures of Income Inequality from Survey Data: An Application of the Estimating Equations. Survey Methodology 21(2): 137-145.
- Kovacević, Milorad S., David A. Binder (1997). Variance Estimation for Measures of Income Inequality and Polarization – The Estimating Equations Approach. Journal of Official Statistics 13(1): 41-58.